3

# CONTINUOUS NUMERICAL OUTCOMES

## **Contexts**

Our actions are only throws of the dice in the sightless night of chance

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What is this chapter about?

In this chapter we consider the probabilities associated with continuous random variables; that is, numerical outcomes that can take any value within a range – for instance, the height of a 20-year-old male, the duration of a call to a call centre, or the depth of the sea at a randomly chosen point.

## Why is it useful?

As we said in the previous chapter, the ability to predict the likelihood of various values is invaluable for planning and decision making. This chapter widens the repertoire of 'patterns' of probabilities or probability distributions. In particular, it includes the most commonly occurring distribution: the normal distribution, which is essential for inferential statistics.

## Where does it fit in?

Much of the subject of Probability is concerned with the probabilities associated with random variables; that is, measurements that take an uncertain numerical value. There are two types of random variable: discrete and continuous. The last chapter (P2) was concerned with discrete random variables and this chapter considers continuous random variables.

## What do I need to know?

You must be at ease with the last two chapters (P1 and P2) and, of course, with the material in *Essential Maths.* 

## **Objectives**

After your work on this chapter you should be:

- able to use a probability density function (pdf) to represent the probabilities associated with a continuous random variable;
- able to calculate the probabilities associated with the uniform distributions;
- familiar with the normal distribution and the standard normal distribution and able to calculate their probabilities using tables or a computer.

After considering discrete random variables in Chapter P2 we now move on to continuous random variables – that is, random variables that can take *any* value within a range.

We start this chapter by introducing the *probability density function* (*pdf*) of a continuous random variable and showing how probabilities can be calculated from it. The normal distribution is particularly important as it is perhaps the most widely used distribution and it plays a key role in Statistics.

## **1** Probability density functions (pdfs)

A continuous random variable is a random variable that can take *any value* within a range, as in the following examples.

- The annual rate of inflation for the UK.
- The time it takes a runner to complete a marathon.
- The time it takes a particular Formula One racing driver to complete a lap at the Monaco Grand Prix.
- Human body temperature when in good health.
- The time between phone calls received at a switchboard.
- The weight of a new-born baby.
- The amount of your electricity bill in the winter quarter.

Notice that continuous random variables are usually measurements as opposed to 'counts'. Contrast these with discrete random variables that can only take *a selection of values* within their range, for instance, the number of cars that pass a motorway check-point in a minute, or the number of students who pass an exam out of a class of 30.

## The probability distribution of a continuous random variable

We already know that the probability distribution of a discrete random variable is a list of all its possible values and their probabilities. For example, a typical probability distribution for a discrete random variable, *X*, might be

x	<b>P</b> (x)	
0	0.01	
1	0.5	
2	0.4	
3	0.07	
4	0.02	

Before you read on, can you see why we can't represent the probabilities of a continuous random variable in the same way?

The reason is that a continuous random variable can take *any value* within its range and so a list of all the possible values would be infinitely long. For instance, the time it takes a runner to complete a marathon can be recorded to any number of decimal places and so, in theory, an infinite number of different times are possible. In practice of course, the time will only be recorded to the nearest tenth or hundredth of a second, and so there will be a very large number rather than an infinite number of possible times.

So, we do not write down the probability distribution of a continuous random variable in the manner shown above because the number of probabilities required would be infinite or, at best, prohibitively large. Also, as the total of the probabilities in a probability distribution has to be 1, most of these probabilities would be extremely small.

To avoid this problem, the probability structure associated with a continuous random variable is represented using a *probability density function* (*pdf*) (sometimes just called the *probability distribution* again). The graph in Figure 3.1 shows the pdf of a random variable, *X*. As you can see, it is just a curve that lies on or above the *x* axis of a graph.

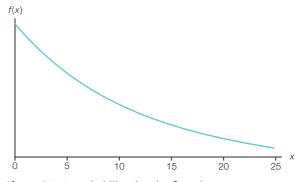
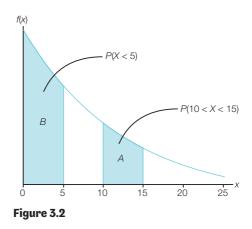


Figure 3.1 A probability density function.

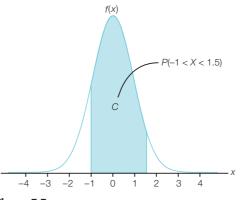
The pdf is chosen so that the area underneath the curve between any two values on the *x* axis is the probability that *X* falls between these two values. For instance, area *A* in the graph in Figure 3.2 shows the probability that *X* lies between 10 and 15, P(10 < X < 15) and area *B* shows the probability that *X* is less than 5; that is, P(X < 5).



It follows that the total area under a pdf curve must be 1 because the probability that *X* takes a value that lies somewhere on the *x* axis is 1.

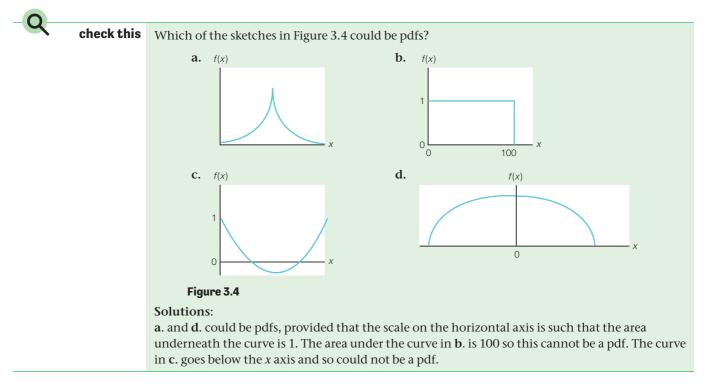
The pdf curve is usually given by a formula in *x* which we will refer to as f(x). For instance, the pdf shown in Figures 3.1 and 3.2 is  $f(x) = 0.1e^{-0.1x}$ , where e is just the number 2.71828... as usual.

As another example, consider the pdf in Figure 3.3. As before, the area under the whole curve is 1 but this time X can take both positive and negative values. Area C is the probability that X lies between -1 and 1.5.



The pdf of a random variable X is a curve, given by the formula f(x), with the following properties:

- (i) it lies on or above the *x* axis,
- (ii) the area under the whole curve is 1, and
- (iii) the area under the curve between any two values *a* and *b* is P(a < X < b).



As a continuous random variable can take an infinite number of possible values, the probability that it takes any individual one is 0. This means that for continuous random variables it doesn't matter whether we talk about, for instance, P(X < 5) or  $P(X \le 5)$ , or about P(-1 < X < 1.5) or  $P(-1 \le X \le 5)$  because the end points of any interval contribute a zero probability.

## Using pdfs for discrete random variables

When a discrete random variable has a huge number of possible values it is impractical to work with the (discrete) probability distribution.

For instance, suppose a large motor manufacturer sells between 10,000 and 12,000 cars inclusive a month. The discrete probability distribution of *X*, the number of cars sold in a month, contains 2001 probabilities, P(X = 10,000), P(X = 10,001), P(X = 10,002) and so on up to P(X = 12,000), and the probabilities which are usually of interest like P(X > 10,500) or P(X < 11,700) are generally the sum of many of these. This is cumbersome and we are not likely to be able to estimate all 2001 probabilities accurately. It is much easier to represent the probabilities approximately using a pdf (maybe like the one in Figure 3.5) and treat car sales as a continuous random variable. Any errors due to the fact that *X* is really discrete are usually small.

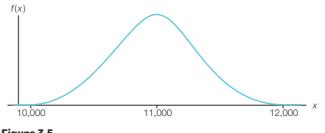


Figure 3.5

## Pdfs and histograms

Suppose *X* is the duration of the working life of a type of electronic chip. The manufacturer takes a sample of 20 such chips and tests their lifetimes. A histogram of the lifetimes of these 20 chips is shown in Figure 3.6.

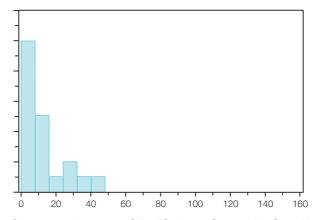
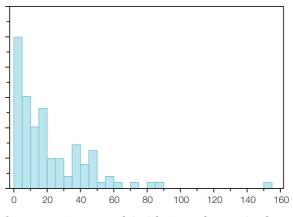


Figure 3.6 Histogram of the lifetimes of a sample of 20 chips.

The distribution is clearly skewed to the right. A larger sample, of 100 chips, gives the histogram in Figure 3.7.



**Figure 3.7** Histogram of the lifetimes of a sample of 100 chips.

This has a similar general shape to the histogram in Figure 3.6 but shows a little more detail because the class width can be narrower as there are more data. A sample of 2000 chips gives the histogram in Figure 3.8, which is even finer.

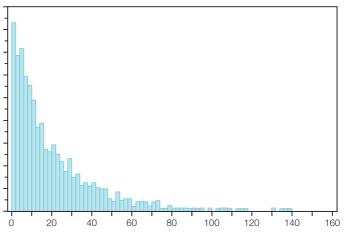
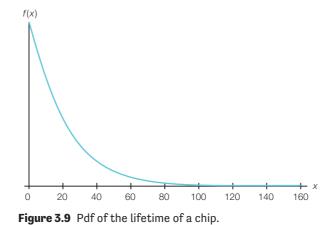


Figure 3.8 Histogram of the lifetimes of a sample of 2000 chips.

We could continue by taking larger and larger samples. As the size of the sample increased the class width could be reduced further and the corresponding histograms would gradually approach the shape of the pdf of the lifetime of a chip, which is shown in Figure 3.9.



So the pdf can be regarded as the histogram of an infinite number of repetitions of the random variable. (We had a similar result for discrete random variables - that the relative frequencies of a large number of values of a discrete random variable approach the probabilities (see p. 295).)

## Drawing pdfs

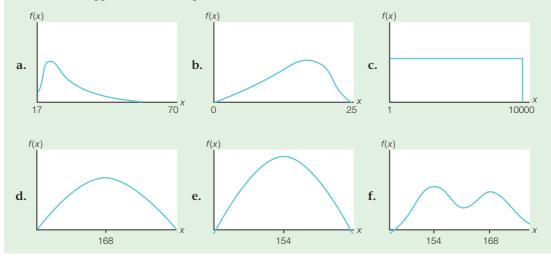
To get a feel for pdfs try sketching a rough pdf for the following random variables. Remember that you are really just drawing a histogram of lots of repetitions of the random variable.

check this Draw a rough graph of a plausible pdf for the following random variables.

- a. The age of an undergraduate student selected at random from your university.
- **b**. The number of minutes to the next bus, when you have just missed one.
- c. The winning ticket in a lottery in which 10,000 tickets are numbered consecutively from 1 to 10,000.
- d. The height of an adult man.
- e. The height of an adult woman.
- **f**. The height of an adult.

#### Solutions:

No exact answers are possible here – it depends what is assumed about the random variable of interest. We suggest the following sketches.



#### check this continued

Q

- a. We have assumed that a student cannot be younger than 17, that the ages between about 18 and 22 are most common and that the higher the age above this, the less common it is.
  - **b**. We have assumed that until a due time of 20 minutes the bus becomes more likely to arrive, and that one is certain to arrive within 25 minutes.
  - **c.** This is a discrete random variable but there are 10,000 possibilities so we can approximate the probability distribution using a pdf. Each ticket number is equally likely so the pdf has equal height throughout the range.
  - **d**. We have assumed that the most likely height of a man is about 168 cm and that the more a height deviates from this the less likely it is.
  - e. We have assumed that the most likely height of a woman is 154 cm and that the more a height deviates from this the less likely it is.
  - **f**. As half of all adults are men and half are women the resulting distribution will be a combination of those in **d**. and **e**. and a pdf with two modes (peaks) will result.

#### Calculating continuous probabilities: the uniform distribution

When all the possible values of a continuous random variable are equally likely to occur we say it has a *uniform* distribution. The uniform distribution is the simplest of a few well-known continuous distributions that occur often. Consider the following example.

An office fire-drill is scheduled for a particular day, and the fire alarm is equally likely to ring at any time between 9am and 5pm. The time the fire alarm starts, measured in minutes after 9am, is therefore a random variable, *X*, which is equally likely to take any value between 0 and 480. A sketch of the pdf of *X*, with a reminder that the total area under a pdf is 1, is shown in Figure 3.10.



Figure 3.10

The pdf is the same height throughout the range of *X* and so the equation of this pdf has the form f(x) = height. The area under the pdf, which must equal 1, is a rectangle with base 480, so  $480 \times \text{height} = 1$ , and it follows that the height must be 1/480. So the pdf is f(x) = 1/480 for any value of *x* between 0 and 480. We write this as

$$f(x) = \frac{1}{480}$$
 for  $0 \le x \le 480$ 

Now we have found the pdf we can use it to calculate probabilities about *X*. For instance, the probability that the fire alarm sounds between 1pm and 2pm, P(240 < X < 300), is the area under the pdf between 240 and 300, which is shown in Figure 3.11.

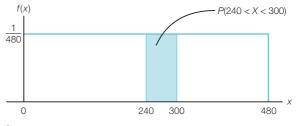
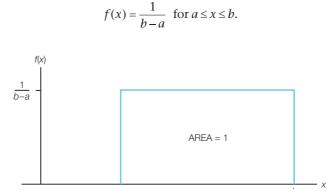


Figure 3.11

This area is a rectangle with base 60 (= 300 – 240) and height  $\frac{1}{480}$ , so  $P(240 < X < 300) = \frac{1}{8}$ . (Some readers may think that this probability was obvious from the start, but we have used this method because it applies to more complicated pdfs as well.)

In general, when a random variable is equally likely to take any value between a and b (where a < b), we say it has a *uniform distribution over the interval from* a *to b*, and because the area under the the pdf between a and b must be 1, as shown in Figure 3.12, its pdf is



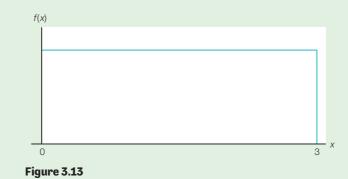


**check this** A local authority is responsible for a stretch of road 3km long through a town. A gas main runs along the length of the road. The gas company has requested permission to dig up the road in one place but has neglected to tell the local authority exactly where.

- **a**. Let *X* be the distance of the roadworks from one end of the road. Sketch the pdf of *X*.
- **b**. What is *f*(*x*)?
- **c.** The stretch of road between 1.5 and 2.75 km from one end goes through the town centre and roadworks there would cause severe disruption. What is the probability that this happens?

#### Solution:

**a**. Assuming that any point along the road from 0 to 3 km from one end is equally likely, we have a uniform distribution over the interval 0 to 3, as shown in Figure 3.13.



- **b**. The base of the rectangle under this pdf is 3, so, as the area under the rectangle must be 1, the pdf is  $f(x) = \frac{1}{3}$  for  $0 \le x \le 3$ .
- c. *P*(1.5 < *X* < 2.75) is the area of the rectangle under the pdf between 1.5 and 2.75. This has height 1/3 and base 1.25, so the probability is 0.4167.

Most pdfs are curves rather than straight lines, so we can't usually use geometry like this to calculate the areas corresponding to probabilities. For some pdfs there is a convenient formula for these areas, and for others specially compiled tables of probabilities are available.

## The expected value and variance of a continuous distribution

In Chapter P2 we introduced the concepts of the expected value (mean) and variance of a discrete random variable. These ideas extend to continuous random variables, although the method of calculation for continuous random variables requires a mathematical method called integration which is not covered by this book. This doesn't matter, however, because there are formulae for the mean and variance of all the well-known distributions, which we will give as we introduce each distribution.

For instance, the mean of a random variable that is uniformly distributed in the interval from *a* to *b*, and so has pdf

 $f(x) = \frac{1}{h-a}$ 

 $\mu = \frac{a+b}{2}$ 

is

and the variance is

 $\sigma^2 = \frac{(b-a)^2}{12}$ 

(The mean may be obvious to you as the distribution is symmetric and  $\frac{a+b}{2}$  is in the middle.) For example, for the fire alarm problem on p. 315, where *X* was equally likely to take any value between 0 and 480, *a* = 0 and *b* = 480. Therefore, the mean is

 $\mu = \frac{0+480}{2} = 240$ 

and the variance is

 $\sigma^2 = \frac{\left(480 - 0\right)^2}{12} = 19,200$ 

**check this** A train is equally likely to arrive at a station at any time between 6.10 pm and 6.40pm.

**a.** What is the pdf of *X*, the number of minutes it arrives past 6pm?

**b.** Calculate the mean and variance of *X*.

#### Solution:

Q

a. X is equally likely to take any value between 10 and 40, so

$$f(x) = \frac{1}{30} \qquad 10 \le x \le 40$$

**b.** The mean of X is

$$\frac{a+b}{2} = \frac{10+40}{2} = 25$$

The variance of *X* is

$$\frac{(b-a)^2}{12} = \frac{(40-10)^2}{12} = 75$$

As usual, the mean and variance are summary measures of the centre and spread of the distribution, respectively. The graph of the pdf gives a useful way of thinking of the mean. Imagine that the pdf and the area under it has been cut out in wood. The point of the *x* axis under which you could place a pivot, to make a balanced see-saw, is the mean. Ρ

We can place the same interpretation on the variance as we did for discrete distributions in Chapter P2, Section 3; that is, when *k* is any positive number greater than 1:

The probability that a random variable X lies within  $k\sigma$  of its mean,  $\mu$ , is at least  $1 - \frac{1}{k^2}$ 

For example, suppose the average journey time by train from Norwich to Nottingham has a mean of 2.8 hours and a variance of 0.16 hours. The standard deviation is  $\sqrt{0.16} = 0.4$ , so we can instantly make the statement that the probability that the journey lasts between 2.8 – (2 × 0.4) and 2.8 + (2 × 0.4) hours (that is, P(2 < X < 3.6)) is at least

$$1 - \frac{1}{2^2} = 0.75$$

or that the probability that the journey lasts between  $2.8 - (3 \times 0.4)$  and  $2.8 + (3 \times 0.4)$  hours (that is, P(1.6 < X < 4)) is at least

$$1 - \frac{1}{3^2} = 0.8889$$

Restating these results in general terms, the probability that a value is within twice the variance of the mean is 0.75 and the probability that a value is within three times the variance of the mean is 0.8889. This will be true for any pdf, whatever the shape of the curve.

Whilst these statements may seem rather vague, they do give some feel for the likely spread of values for the random variable.

#### **Continuous distributions**

The pdf of a continuous random variable, X, is a function f(x) such that

- (i) its curve lies on or above the *x* axis
- (ii) the area under the whole curve is 1

(iii) the area under the curve between *a* and *b* is P(a < X < b)

#### The uniform distribution

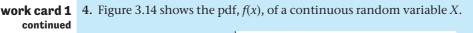
The pdf 
$$f(x) = \frac{1}{b-a}$$
  $(a \le x \le b)$  has mean  $\mu = \frac{b+a}{2}$  and variance  $\sigma^2 = \frac{(b-a)^2}{12}$ 

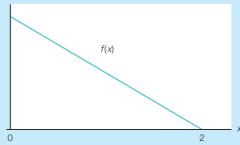
**work card 1** 1. Which of the following would you treat as continuous random variables?

- **a.** The number of goals scored in a football match by a particular team.
  - **b.** The proportion of a cinema audience who smoke cigarettes.
  - c. The number of washing machines sold in a day by a small shop.
  - **d**. The number of washing machines sold in a day by a large retail firm.
  - e. The amount owed by an individual to the Tax office, or owed by the Tax office to her, at the end of a year.
  - f. The *average* number of goals scored in a match by a particular football team during the whole season.
- 2. Where possible, for each of the random variables in question 1, draw a possible probability density function (pdf).
- **3.** An Anglia Railways train is due to arrive at 5.30 pm, but in practice is equally likely to arrive at any time between *2 minutes early and 30 minutes late*. Let the time of arrival (expressed as minutes from due time) be *X*.

Sketch the pdf, f(x), of the random variable *X*. Shade the areas on your graph corresponding to the following probabilities and calculate these.

- **a.** The probability that the train arrives before 5.40 pm.
- **b.** The probability that *X* is less than or equal to 10.
- c. The probability that the train is late, but less than 16 minutes late.





#### Figure 3.14

- **a.** What is the height of the pdf at x = 0?
- **b**. Shade the area corresponding to P(X > 1). Calculate this probability.
- 5. What are the mean and variance of the arrival time of the train in question 3?

#### **Solutions:**

- 1. a. Discrete, b. continuous, c. discrete, d. discrete, but we would model it using a pdf as there are a large number of possible values, e. continuous, f. continuous (notice that, although the number of goals in each match is discrete, the average is continuous).
- **2. b.** The *x* axis must have a range between 0 and 1. We would expect a peak around the proportion of all cinema-goers who smoke.
  - d. Again a peak at the average number sold, positive numbers on the *x* axis only.
  - **e.** Both positive and negative amounts on the *x* axis. Presumably the pdf would tail off for very large positive and very large negative values.
  - **f.** *x* from 0 to 20, say?
- 3. This is a uniform distribution between -2 and 30 (minutes from due time):  $f(x) = \frac{1}{32}$ .
  - **a.** The rectangle under  $f(x) = \frac{1}{32}$  between -2 and 10 with an area of  $\frac{12}{32}$ .
  - **b.** The answer is the same as **a**.
  - **c.** The rectangle under f(x) between 0 and 16 with an area of  $\frac{16}{32}$ .
- 4. a. The triangular area under the pdf must be equal to 1, so the height of the triangle is 1.
  - **b.** Shade the area under f(x) between 1 and 2. This is a triangle with base 1. Its height will be 0.5, as the ratio between the sides is the same as that of the big triangle, and so P(1 < X < 2) = 0.25.
- 5. The train time has a uniform distribution between -2 and 30 so a = -2 and b = 30 and, using the formula for the mean and variance of a uniform distribution, we have

$$\mu = \frac{a+b}{2} = \frac{-2+30}{2} = 14$$

and

$$\sigma^2 = \frac{(b-a)^2}{12} = \frac{(30+2)^2}{12} = 85.33\dot{3}$$

assessment 1	1. Which of the following would you treat as continuous random variables? Explain your
	answers.
	a. The number of people who arrive for a particular charter flight.
	<b>b.</b> The number of mm of rain that falls on London on a particular day.
	c. The number of staff in a small office who are off sick on a particular day.
	d. The proportion of phone calls from a particular pay-phone on a particular day that take
	longer than two minutes.
	e. The duration of a phone call from a pay-phone.
continued overleaf	<b>f.</b> The number of telephone units on a telephone card that a particular call uses.

assessment	
continue	assumptions you make.
	a. The percentage mark gained by a less able individual in an exam.
	<b>b.</b> The percentage mark gained by an able individual in an exam.
	c. The proportion of able students who pass the exam.
	d. The proportion of less able students who pass the exam.
	e. The time it takes you to travel to university on a typical day.
	f. The percentage return over a month on a particular equity investment.
	<b>3.</b> A random variable, <i>Y</i> , can take values between 0 and 0.5 and has the pdf $f(y) = 8y$ .
	<b>a.</b> Calculate the probability that <i>Y</i> is larger than 0.25.
	<b>b.</b> Calculate the probability that <i>Y</i> is smaller than 0.1.
	<b>c.</b> Calculate the probability that <i>Y</i> lies between 0.1 and 0.25.
	4. A financial analyst predicts that the return on Whizzco ordinary shares over the next
	12-month period is equally likely to be any percentage between $-2\%$ and $4\%$ . Assuming that the analyst is correct,
	<b>a.</b> Draw the probability density function of the return, X, and write down its formula, $f(x)$ .
	<b>b.</b> Calculate the probability that the share has a positive return.
	c. Write down the mean and variance of the share's return.

## 2 Introducing the normal distribution

The normal distribution is probably the most widely used and frequently occurring probability distribution. We will explain later why this is so. (By the way, 'normal' is the name of the distribution. It does not mean that other distributions are *ab*normal in any way!)

The pdf of a normal random variable, *X*, is 'bell-shaped' and symmetric about a value,  $\mu$  ('mu'), as shown in Figure 3.15. The curve never actually touches the *x* axis, so any value of *X* is possible, although values of *X* a long way from  $\mu$  are very unlikely.

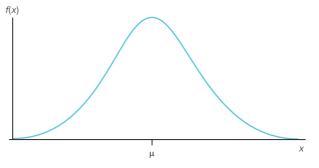


Figure 3.15 The pdf of a normal random variable.

The formula for the pdf of the normal distribution is

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$$

We will not use this again so don't worry if it looks rather formidable, we just ask you to notice that it involves two constants,  $\sigma$  and  $\mu$  (sigma and mu). There are a whole family of normal curves, each one specified by a particular pair of values for  $\sigma$  and  $\mu$ .

Both  $\mu$  and  $\sigma$  have a 'nice' interpretation. We have already said that the pdf is symmetric about  $\mu$ , so it is no surprise that  $\mu$  is the mean of the distribution. The other constant,  $\sigma$ , dictates how spread out and flat the 'bell-shape' is, and in fact  $\sigma^2$  is the variance of the normal distribution. As an illustration, Figure 3.16 shows the following normal pdfs:

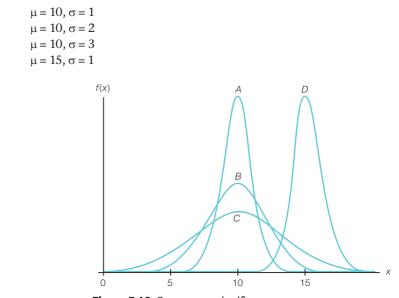


Figure 3.16 Some normal pdfs.

Pdfs *A*, *B* and *C* all have the mean 10, so they are all centred at x = 10. Of these three curves, *C* has the largest variance and so is the most 'spread out'. Curve *B* has a smaller variance and so is less spread out, and curve *A* has the smallest variance and so is the most 'squeezed in'. Curves *A* and *D* have the same variance, so they have exactly the same shape, but they have different means so they are centred at x = 10 and x = 15 respectively.

#### Some notation

A

В

C

D

As the normal distribution is entirely specified by its mean and variance, a shorthand notation has been developed.  $X \sim N(\mu, \sigma^2)$  means that the random variable *X* has a normal distribution with mean  $\mu$  and variance  $\sigma^2$ . So, for instance, the curve shown in *A* above is the pdf of a *N*(10,1) distribution, the curve in *B* is the pdf of a *N*(10,4) distribution and so on.

### The standard normal distribution

The normal distribution with mean  $\mu = 0$  and variance  $\sigma^2 = 1$ , is called the *standard normal distribution*. The letter *Z* is usually used for a random variable that has this distribution. A graph of the standard normal pdf,  $\phi(z)$ , is shown in Figure 3.17. Notice that most of the area under the standard normal curve lies between -3 and +3.

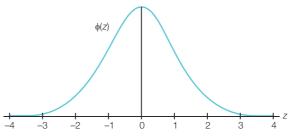


Figure 3.17 The pdf of the standard normal distribution.

## When does the normal distribution arise?

Because the normal pdf peaks at the mean and 'tails off' towards the extremes, the normal distribution provides a good approximation for many naturally occurring random variables. However, the normal distribution occurs even more widely for the following reasons.

1. The total (and also the average) of a large number of random variables that have the same probability distribution has an approximately normal distribution. For instance, if the amount taken by a shop in a day has a particular (maybe unknown) probability distribution, the total of 100 days' takings is the sum of 100 identically distributed random variables and so it will (approximately) have a normal distribution.

Many random variables are normal because of this. For example, the amount of rainfall that falls during a month is the total of the amounts of rainfall that have fallen each day or each hour of the month and so is likely to have a normal distribution. In the same way the average or total of a large sample will usually have a normal distribution. This will be important when we consider populations and samples in the *Statistics* part of this book.

**2**. The normal distribution provides approximate probabilities for a distribution called the binomial distribution, which we consider in P4, section 1.

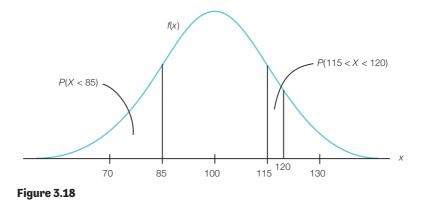
## Calculating probabilities: the standard normal distribution

The normal distribution is continuous and so its probabilities are calculated by obtaining the areas under the pdf curve.

For instance, suppose individuals' IQ scores, *X*, have a normal distribution with mean  $\mu$  = 100 and standard deviation  $\sigma$  = 15. As the standard deviation is 15 we can write this

 $X \sim N(100, 15^2)$  or equivalently  $X \sim N(100, 225)$ .

Figure 3.18 shows the areas under the pdf that correspond to P(X < 85) and P(115 < X < 120).



Unfortunately, there are no 'nice' formulae for calculating such areas so we have to look them up in tables of normal probabilities or use statistical software.

#### Using tables to calculate standard normal probabilities

We will start by evaluating probabilities for a *standard* normal random variable. This may not seem very useful, but in fact such probabilities form the basis of the calculation of *all* normal probabilities, as we shall see in the next section.

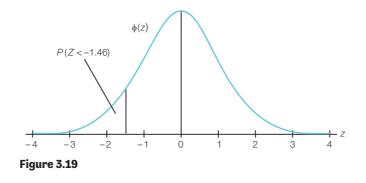
Standard normal probabilities of the form P(Z < a) are listed in a table of *cumulative normal probabilities* such as Table II at the end of this book (pp. 688–689). A section of Table II is shown opposite.

#### Section of Table II

Cumulative standard normal probabilities P(Z < a) where  $Z \sim N(0,1)$ .

а	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
-3.4										
-3.3										
-3.2										
-3.1										
÷										
-1.5	0.0668	0.0655	0.0643	0.0630	0.0618	0.0606	0.0594	0.0582	0.0571	0.0559
-1.4	0.0808	0.0793	0.0778	0.0764	0.0749	0.0735	0.0721	0.0708	0.0694	0.0681
-1.3										
÷										
0.0										
0.1										
0.2										
÷										
1.6	0.9452	0.9463	0.9474	0.9484	0.9495	0.9505	0.9515	0.9525	0.9535	0.9545
:										
1.9	0.9713	0.9719	0.9726	0.9732	0.9738	0.9744	0.9750	0.9756	0.9761	0.9767
2.0	0.9772	0.9778	0.9783	0.9788	0.9793	0.9798	0.9803	0.9808	0.9812	0.9817
÷										
3.3										
3.4										

For example, the probability that *Z* is less than -1.46 (that is, P(Z < -1.46)) shown in Figure 3.19, can be found by looking at the row of the table labelled -1.4 and the column headed 0.06 and is 0.0721.

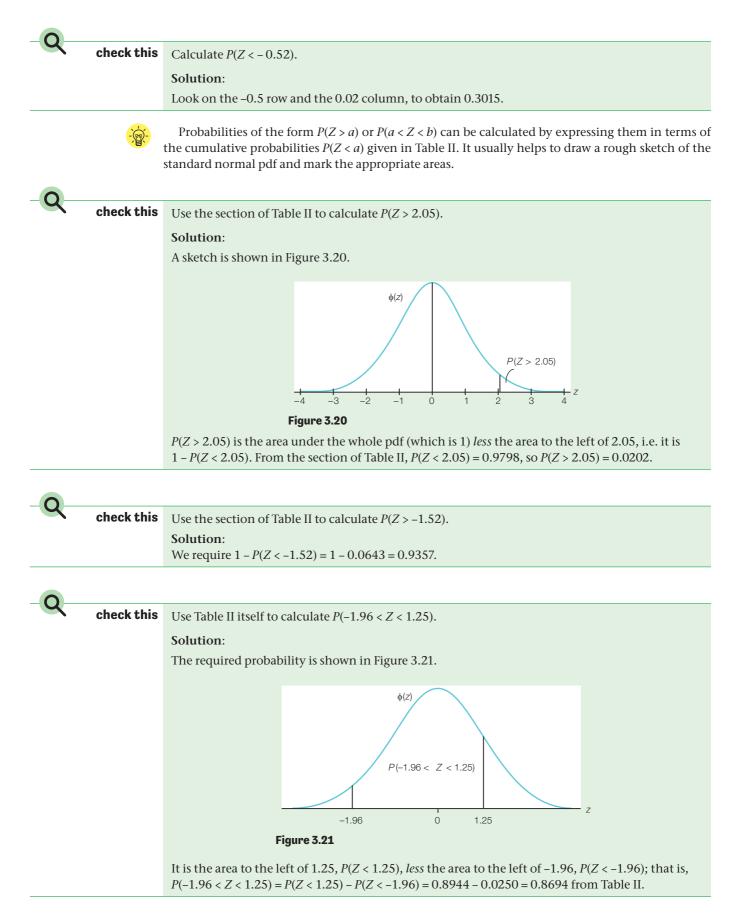


**check this** What is P(Z < 1.96)?

#### Solution:

Look on the 1.9 row and the 0.06 column of the table above to obtain 0.9750.

Now turn to Table II at the end of the book (pp. 688-689).



## Calculating percentage points

Sometimes we will know a cumulative probability but require the corresponding value of *Z*. For instance, suppose we need to find the value *a* such that P(Z < a) = 0.95, as shown in Figure 3.22.

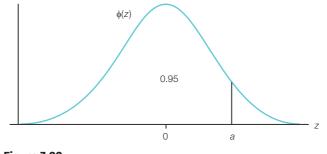


Figure 3.22

There are two ways to find *a*. First, we can use the table of cumulative normal probabilities, Table II, *in reverse*. The body of the table contains P(Z < a) so we need to find 0.95 in the body of the table and then look at the margins to see which value of *a* this corresponds to. On inspection, 0.95 does not explicitly appear in the body of the table – the entry for 1.64 is 0.9495 and the entry for 1.65 is 0.9505. It therefore seems a reasonable guess to suppose that the value we seek lies midway between 1.64 and 1.65. We conclude that P(Z < 1.645) = 0.95, although using the mid-point is an approximation.

The second way avoids this approximation and is the recommended method. Tables of *inverse normal probabilities* or *normal percentage points*, such as the one in Table III (pp. 690–691), give the value of *a* that corresponds to a particular cumulative probability, *p*. An extract from Table III is given below.

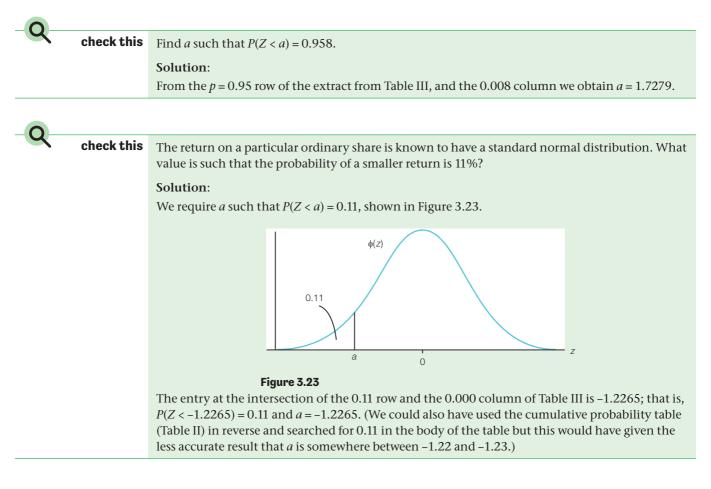
#### Section of Table III

Percentage points of the standard normal distribution The table gives values of *a*, where P(Z < a) = p

p	0.000	0.001	0.002	0.003	0.004	0.005	0.006	0.007	0.008	0.009
0.00		-3.0902	-2.8782	-2.7478	-2.6521	-2.5758	-2.5121	-2.4573	-2.4093	-2.3656
0.01	÷	÷	÷	÷	÷	÷	÷	÷	÷	÷
0.02										
0.03										
÷	÷	÷	÷	÷	÷	÷	÷	÷	÷	÷
0.10	-1.2816	-1.2759	-1.2702	-1.2646	-1.2591	-1.2536	-1.2481	-1.2426	-1.2372	-1.2319
÷	÷	÷	÷	÷	÷	÷	÷	÷	÷	÷
0.95	1.6449	1.6546	1.6646	1.6747	1.6849	1.6954	1.7060	1.7169	1.7279	1.7392
0.96	÷	÷	÷	÷	÷	÷	÷	÷	÷	:
0.97										
0.98	÷	÷	÷	÷	÷	÷	÷	÷	÷	:
0.99	2.3263	2.3656	2.4089	2.4573	2.5121	2.5758	2.6521	2.7478	2.8782	3.0902

To find *a* such that P(Z < a) = 0.95 using this table, we look in the row for 0.95, and, as the third decimal place of 0.95 is zero, the column corresponding to 0.000. This gives *a* = 1.6449, so we conclude that P(Z < 1.6449) = 0.95.

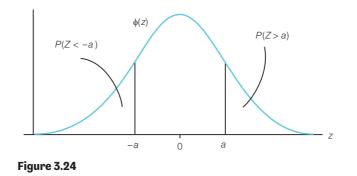
The value of a random variable that corresponds to a cumulative probability of *p* is called the 100*p*th *percentage point* or the 100*p*th *percentile*. So we have just found that the 95th percentage point (or 95th percentile) of the standard normal distribution is 1.6449. We will not use this terminology very much but you may well encounter it elsewhere.



## A dodge for calculating normal probabilities

There is a 'dodge' for calculating normal probabilities that is worth pointing out because it can save time.

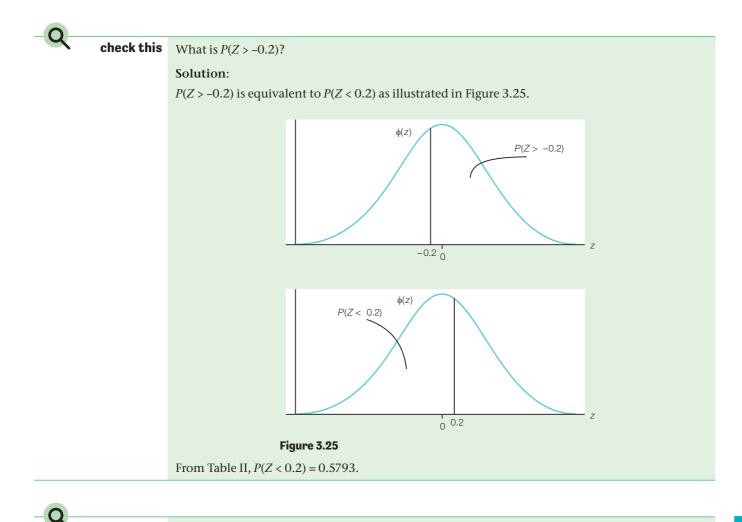
As the standard normal distribution is symmetric about 0, the area under the pdf in the left-hand 'tail', P(Z < -a), is the same as the area under the pdf in the right-hand tail, P(Z > a), as shown in Figure 3.24.



**check this** What is P(Z > 1.3)?

#### Solution:

This is the area under the normal curve to the *right* of 1.3 so we cannot look it up in Table II immediately; we would have to calculate 1 - P(Z < 1.3). However, it is quicker to say that P(Z > 1.3) = P(Z < -1.3) which, from Table II, is 0.0968.

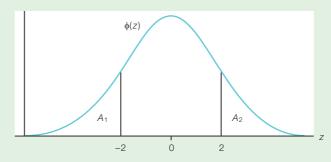


check this

The deviation in mm between the diameter of a manufactured pipe and the specified diameter can be positive or negative (or zero) and has a standard normal distribution. To be usable the diameter of a manufactured pipe must not differ from the specified diameter by more than 2mm. What is the probability of an *un*usable pipe?

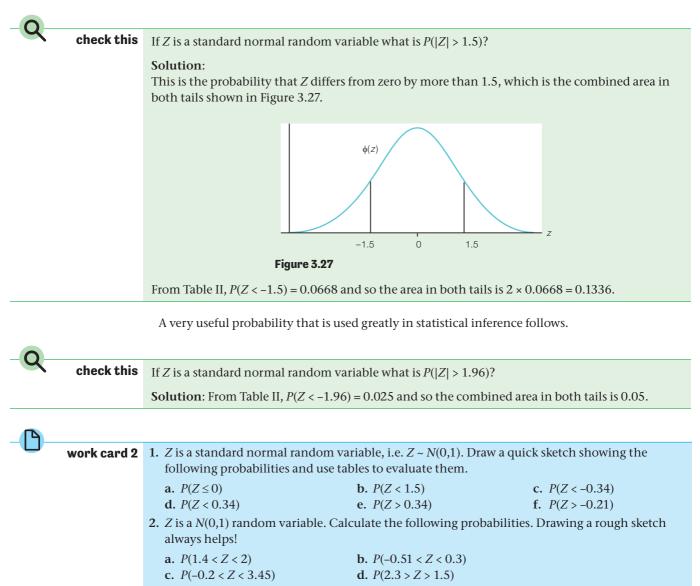
#### Solution:

We need the total of areas  $A_1$  (diameter too small) and  $A_2$  (diameter too large) shown in Figure 3.26 (not to scale).



#### Figure 3.26

We could calculate each of these areas separately and then add them up. However, we know that P(Z < -2) and P(Z > 2) are the same, so we need only calculate one and then double the result.  $A_1 = P(Z < -2)$  is obtained directly from Table II and is 0.0228. So the total area in both 'tails' is  $2 \times 0.0228 = 0.0456$ . We conclude that just under 5% of the pipes are not usable. In the last example it would have been easier to talk about the probability that the *absolute value* of the deviation in the diameter of the pipe, *Z*, is more than 2, written P(|Z| > 2). The absolute value of a number is the amount by which the number differs from zero; that is, its value ignoring any minus sign. For instance, the absolute value of -3, written |-3| is 3 and |-5.2| = 5.2 and so on. So P(|Z| > 2) means the probability that *Z* differs from zero by more than 2, which points back to the combined areas in the tail to the left of -2 and the tail to the right of 2.



- **3.** *Z* is a standard normal random variable.
  - **a.** Find a value *a* such that there is a  $2\frac{1}{2}$ % probability that *Z* is larger than *a*.
  - **b.** Find the value *a* such that the probability of a smaller value is 0.01.
  - **c.** Find *a* such that P(Z < a) = 0.01.
  - **d.** Find *a* such that P(Z > a) is 0.63.

#### 4. Assuming that *Z* is a standard normal random variable, answer these *without* using tables.

- **a.** If P(Z > 1.35) = 0.0885 what is P(Z < -1.35)?
- **b.** If P(Z > 1.36) = 0.0869 what is P(Z > -1.36)?
- c. If P(Z > 1.9) = 0.0287 what is P(-1.9 < Z < 1.9)?
- **d.** If P(Z < -2.84) = 0.0023 what is P(|Z| > 2.84)?
- e. Find two values such that there is a 95% chance that *Z* lies between them, and 0 lies in the middle of them.

work card 2	Solutions:
continued	<ol> <li>a. You can do this one without tables because <i>z</i> = 0 is the half-way point of a symmetric distribution so <i>P</i>(<i>Z</i> ≤ 0) = 0.5.</li> <li>b. <i>P</i>(<i>Z</i> &lt; 1.5) = 0.9332.</li> <li>c. <i>P</i>(<i>Z</i> &lt; -0.34) = 0.3669.</li> </ol>
	<b>d.</b> $P(Z < 0.34) = 0.6331$ . <b>e.</b> $P(Z > 0.34) = 1 - P(Z < 0.34) = 1 - 0.6331 = 0.3669$ or use $(Z > 0.34) = P(Z < -0.34) = 0.3669$ from c.
	f. $P(Z > -0.21) = 1 - P(Z < -0.21) = 1 - 0.4168 = 0.5832.$ 2. a. $P(1.4 < Z < 2) = P(Z < 2) - P(Z < 1.4) = 0.9772 - 0.9192 = 0.0580.$ b. $P(-0.51 < Z < 0.3) = P(Z < 0.3) - P(Z < -0.51) = 0.6179 - 0.3050 = 0.3129.$ c. $P(-0.2 < Z < 3.45) = P(Z < 3.45) - P(Z < -0.2) = 0.9997 - 0.4207 = 0.579.$ d. $P(2.3 > Z > 1.5) = P(Z < 2.3) - P(Z < 1.5) = 0.9893 - 0.9332 = 0.0561.$
	<ul> <li>3. Use Table III (or, less accurately, look up the probability in the body of Table II).</li> <li>a. P(Z &lt; a) = 0.975, so look up p = 0.975 in Table III to give a = 1.96.</li> <li>b2.3263.</li> <li>c. Same as b.</li> </ul>
	<ul> <li>d. If P(Z &gt; a) = 0.63 it follows that P(Z &lt; a) = 0.37. Looking up p = 0.37 in Table III gives a = -0.3319.</li> <li>4. a. P(Z &lt; -1.35) is identical to P(Z &gt; 1.35) so it too is 0.0885.</li> <li>b. A graph is recommended here. P(Z &gt; -1.36) = 1 − P(Z &lt; -1.36), and P(Z &lt; -1.36) is the same as P(Z &gt; 1.36), so the answer is 1 − 0.0869 = 0.9131.</li> </ul>
	<ul> <li>c. P(-1.9 &lt; Z &lt; 1.9) is the entire area under the normal pdf less the area in each of the two 'tails', P(Z &lt; -1.9) and P(Z &gt; 1.9). Each tail has the same area of 0.0287 so the answer is 1 - (2 × 0.0287) = 0.9426.</li> <li>d. P( Z  &gt; 2.84) means the probability that Z differs from 0 by more than 2.84 and so is the combined area in both tails, 2 × 0.0023 = 0.0046.</li> </ul>
	e. As 0 must lie in the middle, the lower and upper values must be $-a$ and $+a$ , respectively, so the tail areas to the left of $-a$ and right of $a$ are the same. As we require the area under the pdf between the two values to be 0.95, the area in each tail must be 0.025, so $P(Z < -a) = 0.025$ . Using tables, $a = 1.96$ , so the two values are $-1.96$ and $1.96$ .

Ρ

 $\square$ **assessment 2** 1. Calculate the following probabilities when *Z* is a *N*(0,1) random variable. **a.** *P*(*Z* < 1.32) **b.** *P*(*Z* < -1.32) c. P(Z > 2.25)**d.** P(Z < -4.1). 2. In all the following questions *Z* is a standard normal random variable. a. 0.18 is the probability that a standard normal random variable is less than what value? **b.** Find *a* when *P*(*Z* < *a*) = 0.18. **c.** Find *a* such that P(Z > a) = 0.76. **d.** For what value is the probability that *Z* exceeds it, 6%? 3. When *Z* is a standard normal random variable P(Z < -1.5) = 0.0668. Deduce the following without using tables. **a.** P(Z > 1.5)**b.** P(Z > -1.5)c. P(-1.5 < Z < 1.5)**d.** P(|Z| > 1.5)

## 3 Calculating normal probabilities

We will now see that *any* normal probability can be expressed in terms of a standard normal probability.

## Standardising

Any normal random variable, *X*, that has mean  $\mu$  and variance  $\sigma^2$  can be *standardised* as follows. Take the variable *X*, and

- (i) subtract its mean,  $\mu$ , and then
- (ii) divide by its standard deviation,  $\sigma$ .

We will call the result *Z*, so

$$Z = \frac{X - \mu}{\sigma}$$

For example, suppose that *X* is individuals' IQ scores and that it has a normal distribution with mean  $\mu = 100$  and standard deviation  $\sigma = 15$ , i.e.  $X \sim N(100, 15^2)$ . To standardise *X* we subtract  $\mu = 100$  and divide the result by  $\sigma = 15$  to give

$$Z = \frac{X - 100}{15}$$

In this way, every value of X has a corresponding value of Z. For instance, when X = 130

$$Z = \frac{130 - 100}{15} = 2$$

and when X = 90

$$Z = \frac{90 - 100}{15} = -0.67$$

Notice that *Z* is the number of standard deviations that *X* lies from its mean. So, for this example, an IQ of 130 is 2 standard deviations of 15 above the mean 100 and an IQ of 90 is 0.67 standard deviations of 15 below the mean.



**check this** The percentage monthly return on an ordinary share, *X*, has a normal distribution with mean 3 and variance 4. How would you calculate the standardised return?

In a particular month the return is X = 6. Standardise this and say how many standard deviations it is above or below the mean.

#### Solution:

To standardise we subtract the mean and divide by the standard deviation so the standardised return would be

$$Z = \frac{X-3}{2}$$

When X = 6,

$$Z = \frac{6-3}{2} = 1.5$$

and X = 6 is 1.5 standard deviations above the mean.

## The distribution of standardised normal random variables

The reason for standardising a normal random variable in this way is that: a standardised normal random variable

$$Z = \frac{X - \mu}{\sigma}$$

has a standard normal distribution. That is,

$$Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$$

Another way of saying this is that the number of standard deviations that a normal random variable is away from its mean has a standard normal distribution. So if we take any normal random variable, subtract its mean and then divide by its standard deviation, the resulting random variable will have a standard normal distribution. We are going to use this fact to calculate (non-standard) normal probabilities.

Consider the probability that an individual's IQ score is less than 85, P(X < 85). The corresponding area under the normal pdf with mean  $\mu = 100$  and standard deviation  $\sigma = 15$  is shown in Figure 3.28.

### **Calculating probabilities**

f(x)P(X < 85)70 100 85 115 130 Figure 3.28

We cannot use normal tables directly because these give N(0,1) probabilities. Instead, we will convert the statement X < 85 into an equivalent statement that involves the standardised score

$$Z = \frac{X - 100}{15}$$

because we know it has a standard normal distribution. We start with

X < 85

To turn X into Z we must standardise the X, but to ensure that we preserve the meaning of the statement we must treat the other side of the inequality in exactly the same way. (Otherwise we would end up calculating the probability of another statement, not X < 85.) 'Standardising' both sides gives

$$\frac{X-100}{15} < \frac{85-100}{15}$$

The left-hand side is now a standard normal random variable and so we can call it Z, and we have

$$Z < \frac{85-100}{15}$$

which is

Ρ

So, we have established that the statement we started with, X < 85, is equivalent to Z < -1. This means that whenever an IQ score, X, is less than 85 the corresponding standardised score, Z, will be less than -1, and so the probability we are seeking, P(X < 85), is the same as P(Z < -1).

P(Z < -1) is just a standard normal probability and so we can look it up in Table II in the usual way, which gives 0.1587. We conclude that P(X < 85) = 0.1587.

This process of rewriting a probability statement about *X* in terms of *Z* is not difficult if you are systematic and write down what you are doing at each stage. We would lay out the working we have just done for P(X < 85) as follows.

**check this** *X* has a normal distribution with mean 100 and standard deviation 15. What is the probability that *X* is less than 85?

Solution:

$$P(X < 85) = P\left(\frac{X - 100}{15} < \frac{85 - 100}{15}\right) = P(Z < -1) = 0.1587 \text{ (from tables)}.$$

Try the following.

**check this** Assuming that an individual's IQ score has a  $N(100,15^2)$  distribution, what is the probability that an individual's IQ score is more than 125?

**Solution**: We require

$$P(X > 125) = P\left(\frac{X - 100}{15} > \frac{125 - 100}{15}\right) = P(Z > 1.67)$$

Table II gives P(Z < 1.67) = 0.9525, so we have P(Z > 1.67) = 1 - P(Z < 1.67) = 1 - 0.9525 = 0.0475.

## -Q-

check this

For each of these write down the equivalent *standard* normal probability. (Don't bother to look up the probability in Table II unless you are really keen!)

- **a**. The IQ of randomly chosen university students, *X*, is normally distributed with mean 115 and standard deviation 10. Consider the probability that a student has an IQ of over 150.
- **b**. The number of people who visit an historic monument in a week is normally distributed with a mean of 10,500 and a standard deviation of 600. Consider the probability that fewer than 9000 people visit in a week.
- **c.** The number of cheques processed by a bank each day is normally distributed with a mean of 30,100 and a standard deviation of 2450. Consider the probability that the bank processes more than 32,000 cheques in a day.

#### Solutions:

**a.** 
$$P(X > 150) = P\left(\frac{X - 115}{10} > \frac{150 - 115}{10}\right) = P(Z > 3.5)$$
  
**b.**  $P(X < 9000) = P\left(\frac{X - 10,500}{600} < \frac{9000 - 10,500}{600}\right) = P(Z < -2.5)$   
**c.**  $P(X > 32,000) = P\left(\frac{X - 30,100}{2450} > \frac{32,000 - 30,100}{2450}\right) = P(Z > 0.78)$ 

## check this

**s** A flight is *due* at Heathrow airport at 1800 hours. Its arrival time has a normal distribution with mean 1810 hours and standard deviation 10 minutes.

- a. What is the probability that the flight arrives before its due time?
- **b.** Passengers must check-in for a connecting flight by 1830 at the latest. What is the probability that passengers from the first flight arrive too late for the connecting flight? (Assume no travelling time from aircraft to check-in.)

#### Solution:

Let the time of arrival, in minutes past 1800, be *X*, so  $X \sim N(10,10^2)$ .

**a**. We require

$$P(X < 0) = P\left(\frac{X - 10}{10} < \frac{0 - 10}{10}\right) = P(Z < -1) = 0.1587$$
  
o. 
$$P(X > 30) = P\left(\frac{X - 10}{10} > \frac{30 - 10}{10}\right) = P(Z > 2) = 0.0228$$

Probabilities like P(a < X < b) can be calculated in the same way. The only difference is that when *X* is standardised similar operations must be applied to both *a* and *b*. That is, a < X < b becomes

$\frac{a-\mu}{-\mu} < $	$\frac{X-\mu}{2}$	$< \frac{D-\mu}{2}$	which is	<i>а</i> -µ	$\leq Z < \frac{b-\mu}{2}$
σ	σ	σ		σ	σ

**ર**—

**check this** Individuals' IQ scores have an  $N(100,15^2)$  distribution. What is the probability that an individual's IQ score is between 91 and 121?

**Solution**: We require *P*(91 < *X* < 121). Standardising gives

$$P\left(\frac{91-100}{15} < \frac{X-100}{15} < \frac{121-100}{15}\right)$$

The middle term is a standardised normal random variable and so we have

$$P\left(\frac{-9}{15} < Z < \frac{21}{15}\right) = P(-0.6 < Z < 1.4) = 0.9192 - 0.2743 = 0.6449$$

## Q

**check this** For each of these write down the equivalent standard normal probability.

- **a**. The length of metallic strips produced by a machine has mean 100 cm and variance 2.25 cm. Only strips with a length between 98 and 103 cm are acceptable. What is the probability that a metallic strip has an acceptable length? You may assume that the length of a strip has a normal distribution.
- **b.** Scores in an exam are adjusted so that they have a normal distribution with an average mark of 56 and a standard deviation of 12. Students gaining between 40 and 70 are considered 'mainstream'. Consider the probability that a student gains a 'mainstream' score.

#### Solutions:

a. *X* ~ *N*(100,2.25) so *P*(98 < *X* < 103)

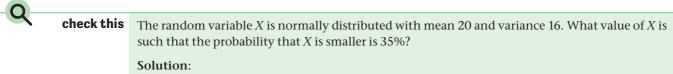
$$= P\left(\frac{98-100}{1.5} < \frac{X-100}{1.5} < \frac{103-100}{1.5}\right) = P(-1.33 < Z < 2)$$

**b.**  $X \sim N(56, 12^2)$  so P(40 < X < 70)

$$= P\left(\frac{40-56}{12} < \frac{X-56}{12} < \frac{70-56}{12}\right) = P(-1.33 < Z < 1.17)$$

## Calculating percentage points

Sometimes you will know a probability but want to calculate the corresponding value of *X*. For instance, you may need to find the value *a* such that P(X < a) = 0.95. Again we can do this by changing the statement *X* < *a* into an equivalent statement about *Z* by standardising *X*.



 $X \sim N(20,16)$  and we require *a* such that P(X < a) = 0.35 as shown in Figure 3.29a.

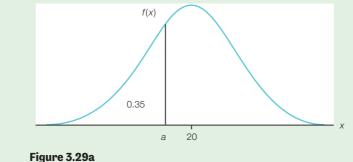


Figure 5.29a

We know P(X < a) = 0.35. Standardising this gives

$$P\left(\frac{X-20}{4} < \frac{a-20}{4}\right) = 0.35$$

 $\frac{X-20}{4}$  is a standardised normal random variable, Z, so our problem is now to find a such that

$$P\left(Z < \frac{a-20}{4}\right) = 0.35$$

as shown in Figure 3.29b.

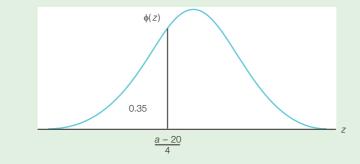


Figure 3.29b

When we look up 0.35 in Table III we obtain -0.3853; that is, P(Z < -0.3853) = 0.35. So, it follows that

$$\frac{a-20}{4} = -0.3853$$

Solving this for *a* gives  $a = (4 \times -0.3853) + 20 = 18.4588$ . So the probability that *X* is less than 18.4588 is 35%.

## -Q check this

According to a survey, dentists' salaries have an average of £48,000 with a standard deviation of £3500. If the salary of a dentist is normally distributed, what salary is such that 20% of dentists have a higher salary?

### Solution:

We require *a* such that P(X > a) = 0.2. We will work in £1000s.

$$X \sim N(48, 3.5^2)$$

so standardising

$$P(X > a) = 0.2$$

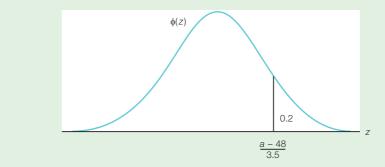
gives

$$P\left(\frac{X-48}{3.5} > \frac{a-48}{3.5}\right) = 0.2$$

that is

$$P\left(Z > \frac{a-48}{3.5}\right) = 0.2$$

as shown in Figure 3.30.





It follows that

$$P\left(Z < \frac{a-48}{3.5}\right) = 0.8$$

and looking up 0.8 in Table III gives

P(Z < 0.8416) = 0.8

So

$$\frac{a-48}{3.5} = 0.8416$$

and solving for *a* gives a = 50.9456. We conclude that 20% of dentists earn more than £50,946.

## check this

The amount a customer spends on a single visit to Rainsburys supermarkets has a normal distribution with mean £75 and standard deviation £21. Rainsburys wish to introduce a minimum amount for which credit cards may be used – an amount which enables 80% of customers to pay by credit card. At what figure should this credit card minimum be set?

#### Solution:

We require *a* such that P(X > a) = 0.8, where  $X \sim N(75,21^2)$ Standardising gives

$$P\left(\frac{X-75}{21} > \frac{a-75}{21}\right) = 0.8$$

which is

$$P\left(Z > \frac{a - 75}{21}\right) = 0.8$$

which is equivalent to

$$P\left(Z < \frac{a - 75}{21}\right) = 0.2$$

Using Table III gives P(Z < -0.8416) = 0.2, so we have

$$\frac{a-75}{21} = -0.8416$$

Solving for *a* gives  $a = (21 \times -0.8416) + 75 = \pounds 57.3264$ . The credit card minimum should be set at £57.33.



## Normal probabilities using software

Most statistical software can calculate cumulative normal probabilities P(X < a) and percentage points (which some software call the *inverse cumulative probabilities*).

Use SPSS's

#### **Transform > Compute Variable**

command and then select the **CDF.NORMAL** function in order to calculate the cumulative normal probability. You will need to specify the values for *a*, the mean and the standard deviation. For instance, when  $X \sim N(75,21^2)$  CDF.NORMAL gives P(X < 50) = 0.1169.

We can find percentage points by selecting the **IDF.NORMAL** function. For the same distribution, we find that the 20th percentage point is 57.3260. This result differs slightly from our earlier manual calculation because the values in Table III are given to only 4 d.p., whereas SPSS's algorithm retains more decimal places in its calculations and so is more accurate.

In Microsoft<sup>®</sup> Excel you need the function NORM.DIST to calculate normal and cumulative normal probabilities, and NORM.INV to calculate the inverse cumulative probability (NORMDIST and NORMINV for Excel 2007 and earlier).

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U	work card 3	<b>1.</b> <i>X</i> is a normal random variable with mean $\mu$ = 5 and variance $\sigma^2$ = 4. Calculate the following						
		probabilities:						
		<b>a.</b> $P(X > 5.7)$ <b>b.</b> $P(X < 3.4)$						
		<b>c.</b> $P(2.8 < X < 5.1)$ <b>d.</b> $P(5.7 < X < 6.8)$						
		<b>2.</b> <i>X</i> is normally distributed with mean 10 and variance 9. Find <i>a</i> such that:						
		<b>a.</b> the probability that <i>X</i> is less than <i>a</i> is 0.51						
		<b>b.</b> $P(X > a) = 0.6$						
		<b>c.</b> $P(X \ge a) = 0.05$						
		<b>d.</b> $P(10 < X < a) = 0.05$						

## 1٩

work card 3 3. The yearly cost of dental claims for the employees of Notooth International is normally distributed with mean  $\mu = \pounds 75$  and standard deviation  $\sigma = \pounds 25$ . Using a computer or otherwise, find:

- a. What proportion of employees can be expected to claim over £120 in a year?
- b. What yearly cost do 30% of employees claim less than?
- 4. Petrol consumption for all types of small car is normally distributed with  $\mu = 30.5$  m.p.g. and  $\sigma = 4.5 \text{ m.p.g.}$

A manufacturer wants to make a car that is more economical than 95% of small cars. What must be its m.p.g? Use a computer if you want.

## **Solutions:**

- 1. a. P(X > 5.7) = P(Z > 0.35) = 0.3632
  - **b.** P(X < 3.4) = P(Z < -0.8) = 0.2119
  - c. P(2.8 < X < 5.1) = P(-1.1 < Z < 0.05) = P(Z < 0.05) P(Z < -1.1) = 0.5199 0.1357 = 0.3842
  - **d.** P(5.7 < X < 6.8) = P(0.35 < Z < 0.9) = P(Z < 0.9) P(Z < 0.35) = 0.8159 0.6368 = 0.1791
- 2. Sketches will help you here.

**a.** P(X < a)

$$= P\left(\frac{X-10}{3} < \frac{a-10}{3}\right) = P\left(Z < \frac{a-10}{3}\right) = 0.51$$
  
Using tables  $P(Z < 0.0251) = 0.51$ , so

$$\frac{a-10}{3} = 0.0251$$
 and  $a = 10.0753$ 

**b.** P(X > a) = 0.6

3

$$P\left(\frac{X-10}{3} > \frac{a-10}{3}\right) = P\left(Z > \frac{a-10}{3}\right) = 0.6$$
  
So  $\frac{a-10}{2} = -0.2533$  and  $a = 9.2401$ 

c.  $P(X \ge a)$ 

$$= P\left(\frac{X-10}{3} \le \frac{a-10}{3}\right) = P\left(Z \le \frac{a-10}{3}\right) = 0.05$$
  
So  $\frac{a-10}{3} = 1.6449$  and  $a = 14.9347$ 

**d.** P(10 < X < a)

$$= P\left(\frac{10-10}{3} < \frac{X-10}{3} < \frac{a-10}{3}\right) = P\left(0 < Z < \frac{a-10}{3}\right)$$
$$= P\left(Z < \frac{a-10}{3}\right) - P(Z < 0) = 0.05$$
So  $P\left(Z < \frac{a-10}{3}\right) - 0.5 = 0.05$  and  $P\left(Z < \frac{a-10}{3}\right) = 0.55$ So  $a - 10 = 0.1257$  and  $a = 10.3771$ 

- **3.** a. P(X > 120) = P(Z > 1.8) = 0.0359, so 3.59%.
  - **b.** We require *a* such that P(X < a) = 0.3, so

$$P\left(Z < \frac{a-75}{25}\right) = 0.3, \ \frac{a-75}{25} = -0.5244 \text{ and } a = 61.89$$

continued overleaf

work card 3

continued

**4**. We require *a* such that P(X < a) = 0.95, so

$$P\left(Z < \frac{a-30.5}{4.5}\right) = 0.95$$
, so  $\frac{a-30.5}{4.5} = 1.6449$ 

and *a* = 37.90205. 95% of small cars do less than 37.9 m.p.g.

**assessment 3** 1. X is a normal random variable with mean  $\mu = -2$  and variance  $\sigma^2 = 0.5$ . Calculate the

- following:
- **a.** *P*(*X* < −3.5) **b.** *P*(*X* > 0)
- c. P(-1.5 < X < -0.8)
- **2.** Calculate the following when  $X \sim N(100,64)$ :
  - **a.** P(X > 120)
  - **b.** *P*(*X* ≤ 99)
  - **c.** P(90 < X < 100)
- 3. *X* has a normal distribution with mean  $\mu = 4$  and standard deviation  $\sigma = 1.5$ . Find *a* when P(4 a < X < 4 + a) = 0.8.
- 4. The duration of a scheduled flight is normally distributed with a mean of 45 minutes and a standard deviation of 2 minutes.
  - a. What is the probability that the flight takes less than 42 minutes?
  - **b.** What is the probability that it takes between 40 and 50 minutes?
  - c. What times do 5% of flights take longer than?
- **5.** For a particular life insurance policy the lifetime of the policyholders follows a normal distribution with mean 72.2 years and standard deviation 4.4 years. One of the options of this policy is that the policyholder receives a payment on their 70th birthday and another payment every 5 years thereafter.
  - a. What percentage of policyholders will receive at least one payment?
  - b. What percentage will receive two or more?



Now, how about trying the set of multiple choice questions for this chapter on the companion website at

www.palgrave.com/companion/Swift-Quantitative-Methods-4

where you can also find other additional resources.